

# Invariant Causal Prediction

## Causal inference in the presence of latent variables

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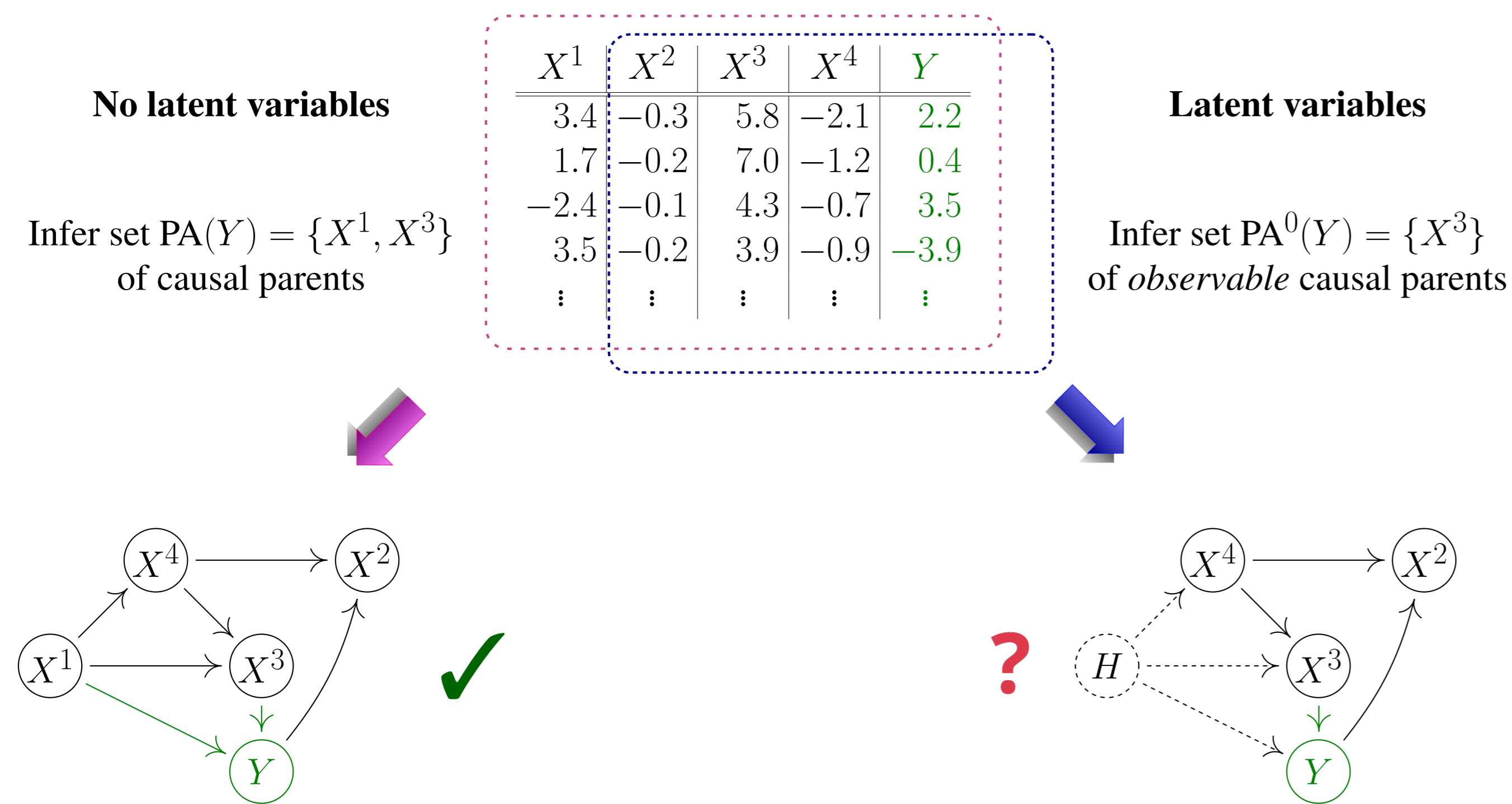
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### Keywords

causal inference invariance latent variables regression analysis hidden Markov models

### Causal discovery and latent variables

Goal: Given a target  $Y$ , find the (observable) causal predictors among  $X^1, \dots, X^d$ .



Problem: In general, inferring  $PA^0(Y)$  from  $P_{(Y, X^2, X^3, X^4)}$  is hard.

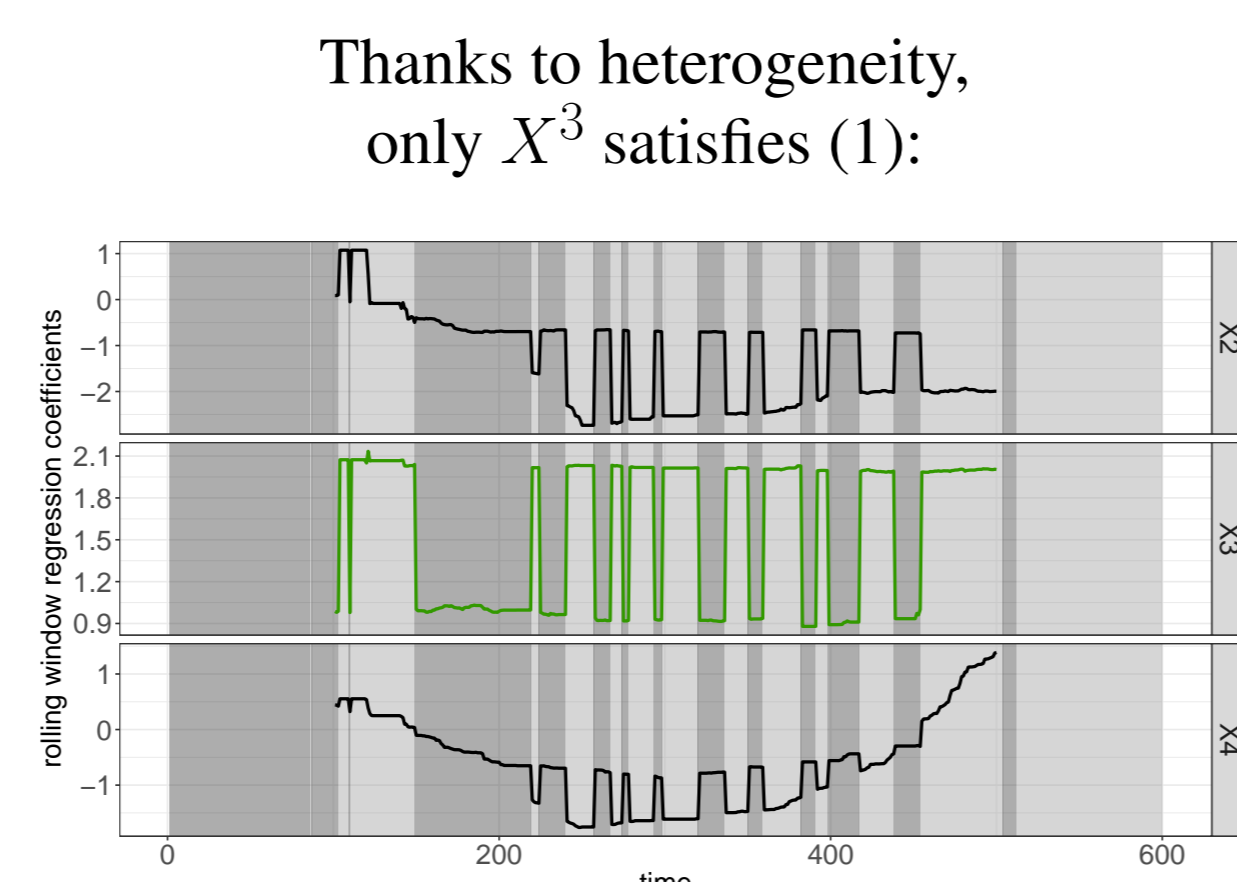
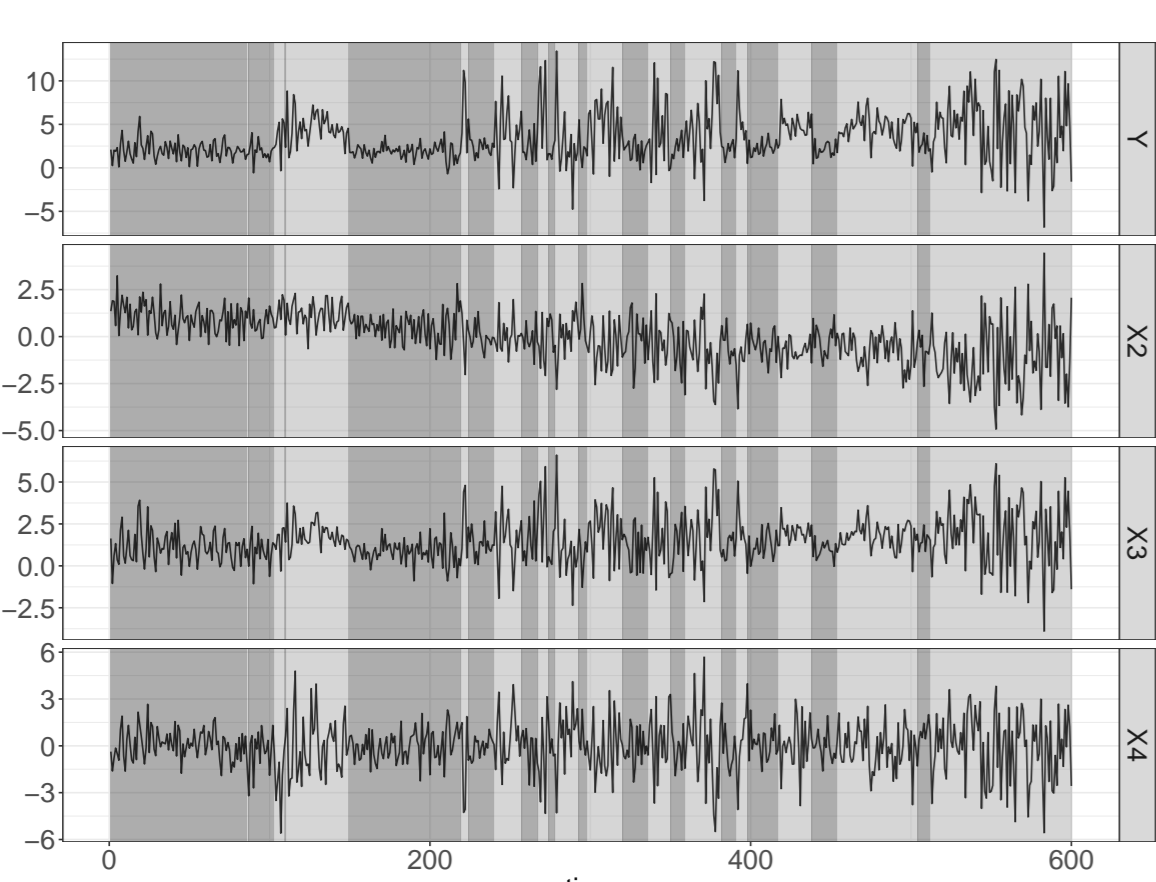
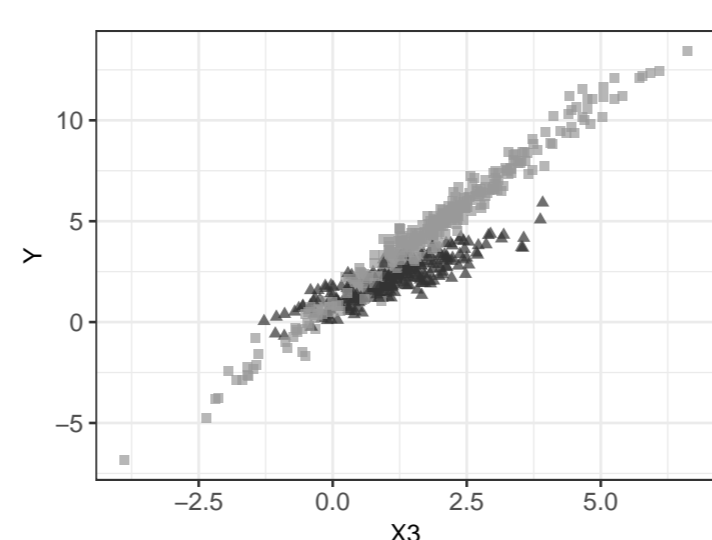
Approach: If  $H$  takes only few discrete values (here  $H \in \{0, 1\}$ ), we can exploit invariances:

### Invariant Causal Prediction

Fundamental assumption:  $(X^3, H) \rightarrow Y$  is time-homogeneous. Then, for all  $t, x$ ,

$$(1) P_{Y_t | (X_t^3=x)} \stackrel{d}{=} \lambda_{xt} P_{Y_t | (X_t^3=x, H_t=0)} + (1 - \lambda_{xt}) P_{Y_t | (X_t^3=x, H_t=1)}$$

independent of  $t$                       independent of  $t$



Key idea:

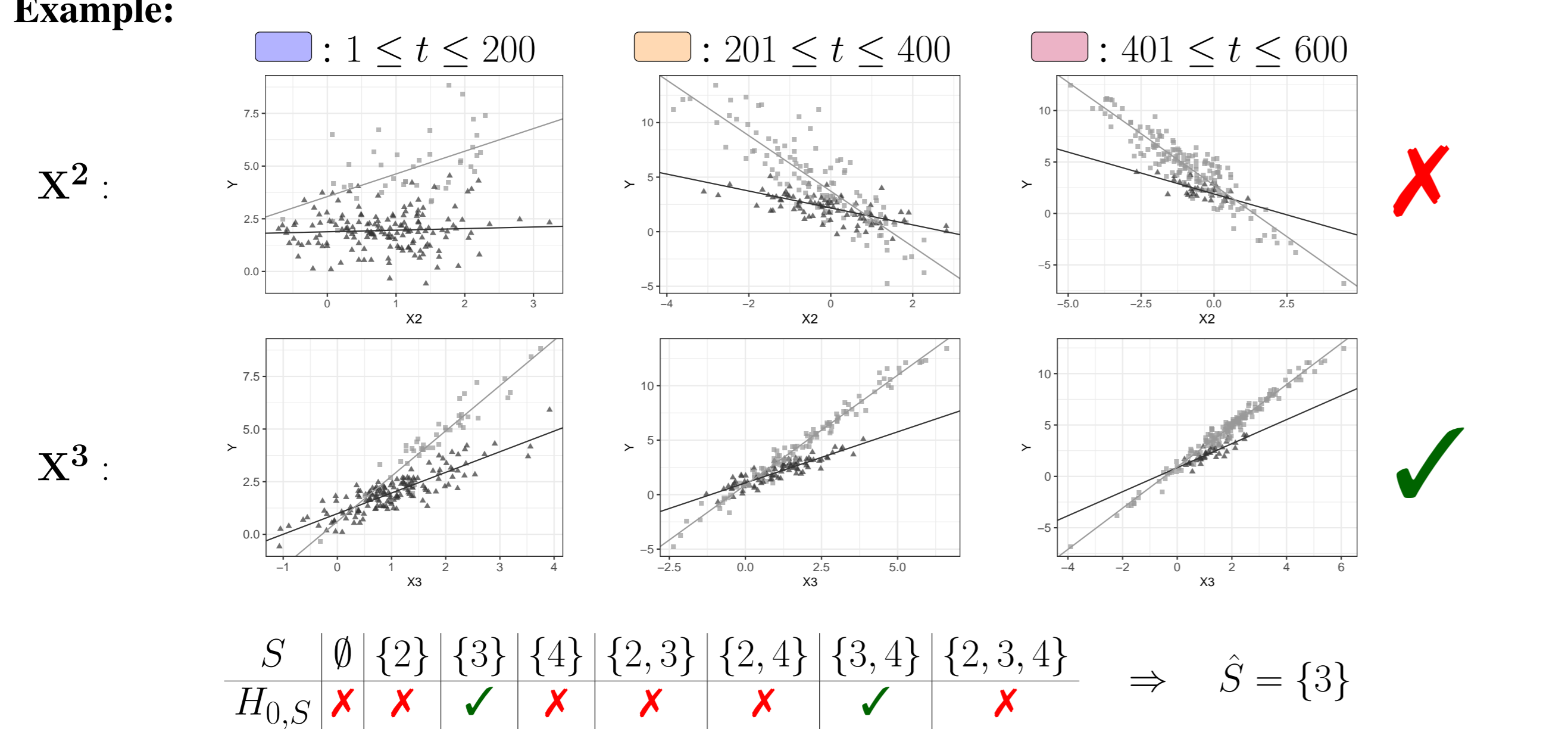
- Use sequential ordering to construct environments  $\square$ ,  $\triangle$  and  $\circ$ .
- For every subset  $S \subseteq \{X^2, X^3, X^4\}$ , test  $H_{0,S} : \{\square, \triangle\} = \{\square, \triangle\} = \{\square, \triangle\}$ .
- Output the estimator  $\hat{S} := \bigcap_{S: H_{0,S} \text{ not rej.}} S$ .

J. Peters, P. Bühlmann, and N. Meinshausen (2016). Causal inference using invariant prediction: identification and confidence intervals. *Journal of the Royal Statistical Society, Series B (with discussion)*, 78(5):947–1012.

Theorem: Assume  $S^*$  is the correct set of observable direct causes. For any test level  $\alpha$  we obtain

$$\lim_{n \rightarrow \infty} \mathbb{P}(\hat{S}_n \subseteq S^*) \geq 1 - \alpha.$$

Example:



### Asymptotically valid hypothesis tests

The false discovery control of our method relies on the asymptotic validity of the tests  $(\varphi_S)_{S \subseteq \{1, \dots, d\}}$  of the hypotheses  $(H_{0,S})_{S \subseteq \{1, \dots, d\}}$ . We assume the true model is: for all  $t, x$ ,

$$(2) P_{Y_t | (X_t^S=x)} = \sum_{j=1}^{\ell} \lambda_{xt}^j \mathcal{N}(x^T \beta_j, \sigma_j^2), \quad \theta = (\beta_1, \sigma_1^2, \dots, \beta_{\ell}, \sigma_{\ell}^2).$$

For every  $S \subseteq \{1, \dots, d\}$ , we construct  $\varphi_S$  by:

- For every environment  $e \in \mathcal{E}$ , fit model (2) to obtain an MLE  $\hat{\theta}_e$  and a Fisher confidence region  $C^\alpha(\hat{\theta}_e)$ .
- Adjust for label permutations, i.e., construct  $C_{\text{adjusted}}^\alpha(\hat{\theta}_e) := \bigcup_{\pi \in \Pi} C^\alpha(\pi(\hat{\theta}_e))$ .
- Reject if  $\bigcap_{e \in \mathcal{E}} C_{\text{adjusted}}^\alpha(\hat{\theta}_e) = \emptyset$ .

Under regularity conditions, these tests are asymptotically valid for any test level  $\alpha$ :

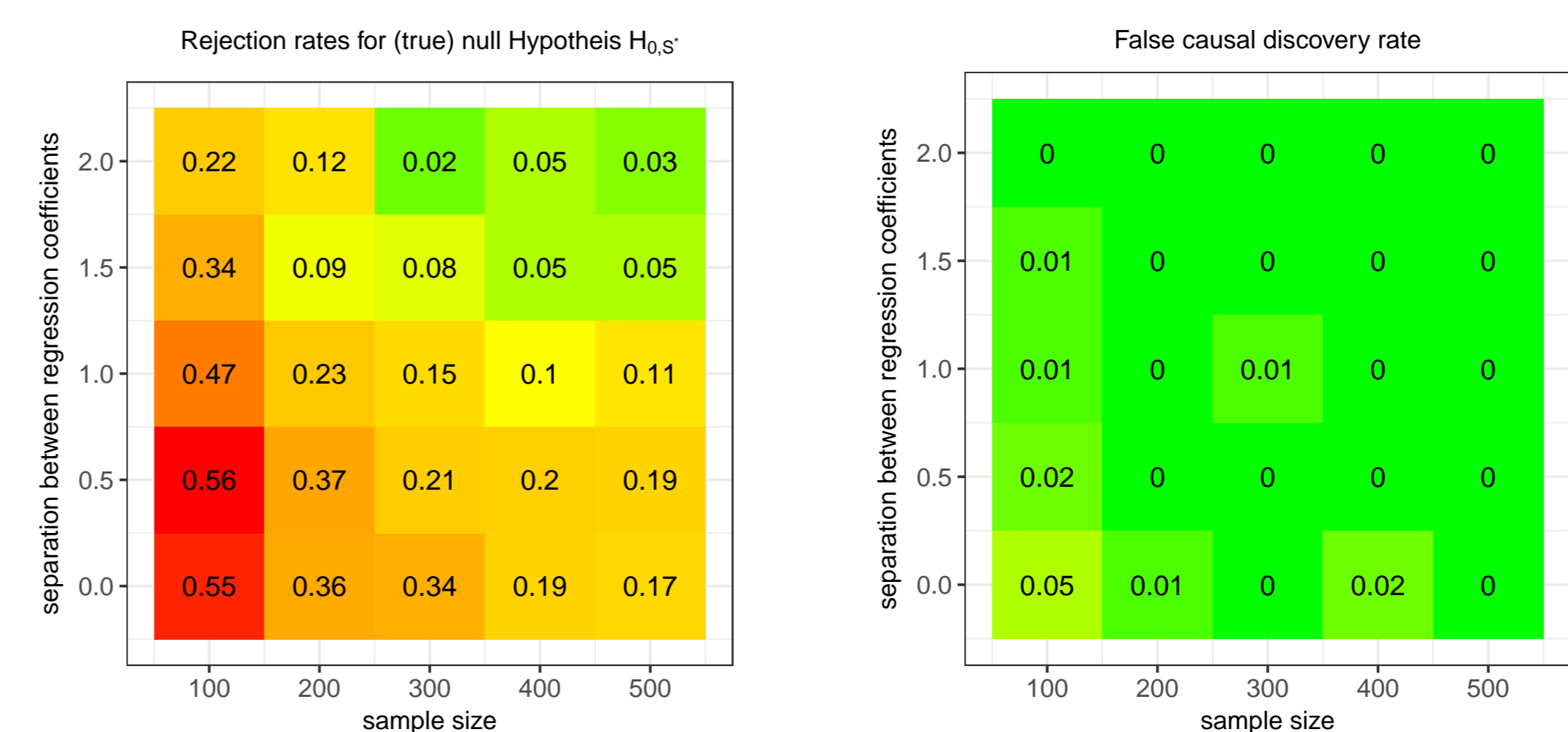
Lemma 1: The MLE is consistent up to label permutations, i.e.,  $\pi_n(\hat{\theta}_n) \rightarrow \theta^0$  as  $n \rightarrow \infty$

Lemma 2: The MLE is asymptotically normal up to label permutations, i.e.,  $\sqrt{n}(\pi_n(\hat{\theta}_n) - \theta^0) \xrightarrow{d} \mathcal{N}(0, \mathcal{I}_0)$  as  $n \rightarrow \infty$

problem:	the likelihood function		
	is non-concave	is unbounded	is based on finite data
solution:	EM-algorithm / non-linear maximization	parameter constraints, e.g., $\min_j \sigma_j^2 \geq c$ or $\sigma_j^2 = \sigma_k^2$	see next section

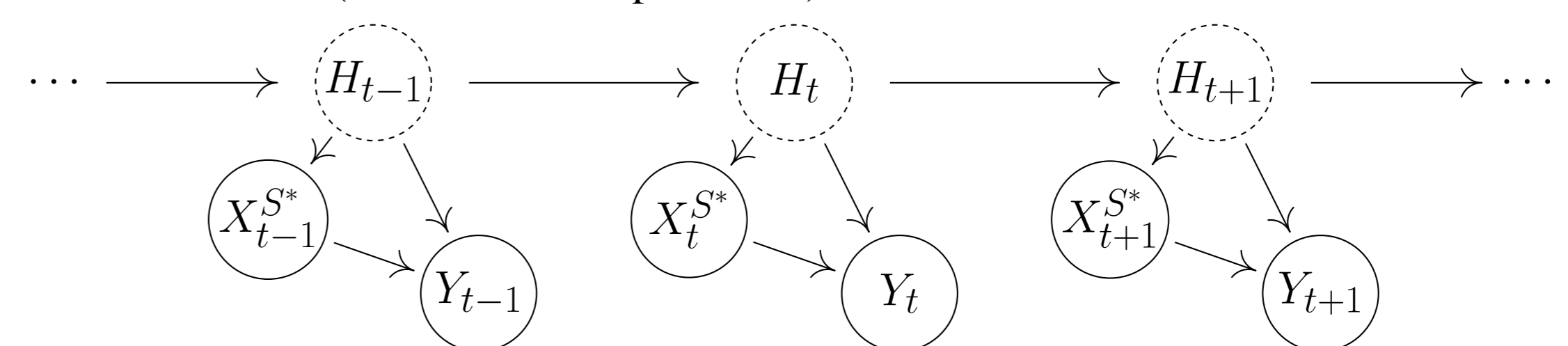
### Robustness for small sample sizes

Even for small sample sizes, the overall method keeps the type I error control.



### Time dependence of latent variables

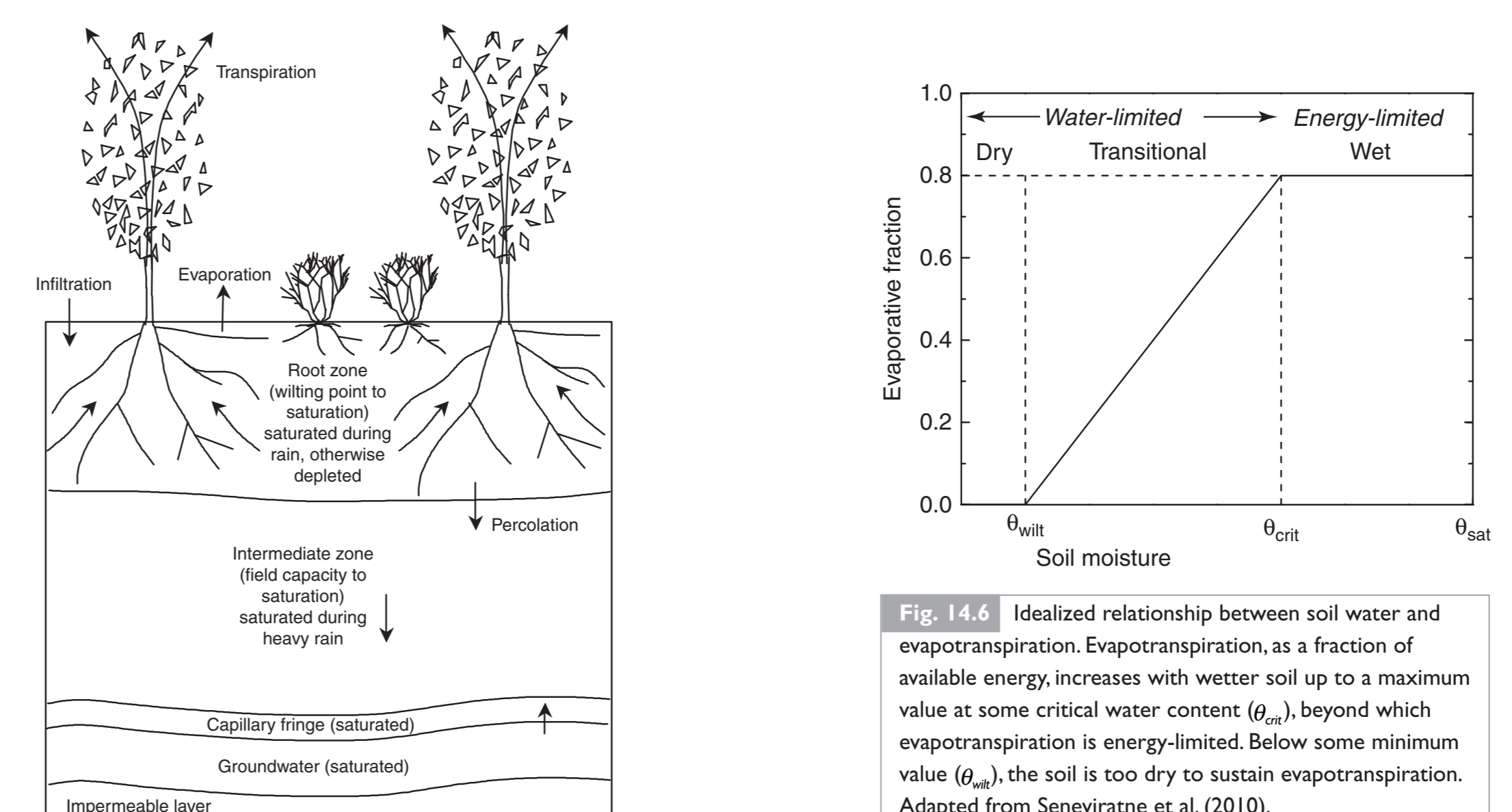
A hidden Markov model (submodel: independence).



### Application in progress

A biogeochemical system:

$Y$ : evapotranspiration,  $X$ : radiation, water pressure deficit, wind speed etc.,  $H$ : soil moisture



G. Bonan (2015). Ecological Climatology: Concepts and Applications. Cambridge University Press.

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